

Hypothesis Testing Workouts: Part III

9. The Exponential Distribution

Greene, Econometric Analysis, 3rd Edition, P. 167

$$n := 20 \quad i := 1 \dots n$$

We have a restricted model of the relationship between income (y) and education (x).

$$f\omega(y, x, \beta) := \frac{1}{\beta + x_i} \cdot e^{\frac{-y_i}{\beta + x_i}}$$

And an unrestricted model

$$f\Omega(y, x, \beta, \rho) := \frac{1}{\Gamma(\rho) \cdot (\beta + x_i)^\rho} \cdot (y_i)^{\rho-1} \cdot e^{\frac{-y_i}{\beta + x_i}}$$

We want to test the hypothesis

$$H_0: \rho := 1$$

$$H_1: \rho <> 1$$

The following sample data has been collected

$$x := \begin{bmatrix} 12 \\ 16 \\ 18 \\ 16 \\ 12 \\ 12 \\ 12 \\ 16 \\ 12 \\ 10 \\ 12 \\ 16 \\ 20 \\ 12 \end{bmatrix} \quad y := \begin{bmatrix} 20.5 \\ 31.5 \\ 47.7 \\ 26.2 \\ 44 \\ 8.28 \\ 30.8 \\ 17.2 \\ 19.9 \\ 9.96 \\ 55.8 \\ 25.2 \\ 29 \end{bmatrix}$$

16	85.5		
10	15.1		
18	28.5	ybar := mean(y)	xbar := mean(x)
16	21.4		
20	17.7	ybar = 31.278	xbar = 14.6
12	6.42	Sy := Stdev(y)	Sx := Stdev(x)
16	84.9	Sy = 22.376	Sx = 3.119

9.a. Likelihood Ratio Test

The unrestricted log likelihood is

$$\ln L(\beta, \rho) := \left[-\rho \cdot \sum_{i=1}^n \ln(\beta + x_i) - n \cdot \ln(\Gamma(\rho)) \right] - \left[\sum_{i=1}^n \left(\frac{y_i}{\beta + x_i} \right) \right] + (\rho - 1) \cdot \left[\sum_{i=1}^n (\ln(y_i)) \right]$$

The maximum occurs where the first derivatives are zero.

$$\frac{d}{d\beta} \left[\left[-\rho \cdot \sum_{i=1}^{20} \ln(\beta + x_i) - n \cdot \ln(\Gamma(\rho)) \right] - \left[\sum_{i=1}^{20} \left(\frac{y_i}{\beta + x_i} \right) \right] + (\rho - 1) \cdot \left[\sum_{i=1}^{20} (\ln(y_i)) \right] \right]$$

$$= -\rho \cdot \sum_{i=1}^n \frac{1}{\beta + x_i} + \left[\sum_{i=1}^n \left[\frac{y_i}{(\beta + x_i)^2} \right] \right]$$

$$\frac{d}{d\rho} \left[\left[-\rho \cdot \sum_{i=1}^{20} \ln(\beta + x_i) - n \cdot \ln(\Gamma(\rho)) \right] - \left[\sum_{i=1}^{20} \left(\frac{y_i}{\beta + x_i} \right) \right] + (\rho - 1) \cdot \left[\sum_{i=1}^{20} (\ln(y_i)) \right] \right]$$

$$= \left(- \sum_{i=1}^{20} \ln(\beta + x_i) - n \cdot \Psi(\rho) \right) + \sum_{i=1}^{20} \ln(y_i)$$

Note: Psi(ρ) is the derivative of $\ln(\Gamma(\rho))$ w.r.t. ρ .

We can use mathCAD's routine's to solve the system of equations. Start by giving the routine

$$\beta := -1 \quad \rho := 1$$

Given

$$\left(- \sum_{i=1}^{20} \ln(\beta + x_i) - n \cdot \text{Psi}(\rho) \right) + \sum_{i=1}^{20} \ln(y_i) = 0$$

$$-\rho \cdot \sum_{i=1}^{20} \frac{1}{\beta + x_i} + \left[\sum_{i=1}^{20} \left[\frac{y_i}{(\beta + x_i)^2} \right] \right] = 0$$

$$\begin{bmatrix} \beta \\ \rho \end{bmatrix} := \text{Find}(\beta, \rho)$$

The unrestricted estimates are

$$\begin{bmatrix} \beta \\ \rho \end{bmatrix} = \begin{bmatrix} -1 \\ 1 \end{bmatrix}$$

The log likelihood at the unrestricted estimates is

$$L\Omega := \ln L\Omega(\beta, \rho)$$

$$L\Omega = -97.033$$

Similarly for the restricted log likelihood. The restricted log likelihood is

$$\ln L\omega(\beta) := -1 \cdot \left(\sum_{i=1}^n \ln(\beta + x_i) \right) - \sum_{i=1}^n \frac{y_i}{\beta + x_i}$$

The maximum occurs at the point where the first derivative w.r.t. β is zero.

$$\frac{d}{d\beta} \left(- \sum_{i=1}^{20} \ln(\beta + x_i) - \sum_{i=1}^{20} \frac{y_i}{\beta + x_i} \right)$$

$$= - \sum_{i=1}^n \frac{1}{\beta + x_i} + \sum_{i=1}^n \frac{y_i}{(\beta + x_i)^2}$$

We can use a MathCAD routine to find the β which makes this f.o.c. zero.

$$\text{root} \left[- \sum_{i=1}^n \frac{1}{\beta + x_i} + \sum_{i=1}^n \frac{y_i}{(\beta + x_i)^2}, \beta \right] = 15.594$$

When $\rho=1$, then the constrained estimate of β is

$$\beta := 15.601$$

The value of the restricted log likelihood is

$$L_{\omega} := \ln L_{\omega}(\beta)$$

$$L_{\omega} = -88.436$$

The likelihood ratio test statistic is

$$LR := -2 \cdot (L_{\omega} - L_{\Omega})$$

$$LR = -17.193$$

This exceeds the critical χ^2 with one degree of freedom at any reasonable level of significance, so reject the null.

9.b. The Wald Test

To do the Wald Test we need the **unrestricted estimates** of the parameters

$$\begin{bmatrix} \beta \\ \rho \end{bmatrix} := \begin{bmatrix} -4.719 \\ 3.151 \end{bmatrix}$$

and the **Information Matrix** evaluated at the unrestricted estimates

$$\frac{d^2}{d\beta^2} \left[\left(-\rho \cdot \sum_{i=1}^{20} \ln(\beta + x_i) - n \cdot \ln(\Gamma(\rho)) \right) - \left[\sum_{i=1}^{20} \left(\frac{y_i}{\beta + x_i} \right) \right] + (\rho - 1) \cdot \left[\sum_{i=1}^{20} (\ln(y_i)) \right] \right]$$

$$= \rho \cdot \sum_{i=1}^n \frac{1}{(\beta + x_i)^2} - 2 \cdot \sum_{i=1}^n \frac{y_i}{(\beta + x_i)^3}$$

$$\frac{d}{d\rho} \left[\left(-\rho \cdot \sum_{i=1}^{20} \ln(\beta + x_i) - n \cdot \ln(\Gamma(\rho)) \right) - \left[\sum_{i=1}^{20} \left(\frac{y_i}{\beta + x_i} \right) \right] + (\rho - 1) \cdot \left[\sum_{i=1}^{20} (\ln(y_i)) \right] \right]$$

$$= \frac{-n \cdot \left[\left[\Gamma(\rho) \cdot \frac{d^2}{d\rho^2} \Gamma(\rho) - \left(\frac{d}{d\rho} \Gamma(\rho) \right)^2 \right] \right]}{\Gamma(\rho)^2} = -7.459$$

$$\frac{d}{d\rho} \left[\frac{d}{d\beta} \left[\left(-\rho \cdot \sum_{i=1}^{20} \ln(\beta + x_i) - n \cdot \ln(\Gamma(\rho)) \right) - \left[\sum_{i=1}^{20} \left(\frac{y_i}{\beta + x_i} \right) \right] + (\rho - 1) \cdot \left[\sum_{i=1}^{20} (\ln(y_i)) \right] \right] \right]$$

$$= - \sum_{i=1}^n \frac{1}{\beta + x_i}$$

$$I := \begin{bmatrix} \rho \cdot \sum_{i=1}^n \frac{1}{(\beta + x_i)^2} - 2 \cdot \sum_{i=1}^n \frac{y_i}{(\beta + x_i)^3} & - \sum_{i=1}^n \frac{1}{\beta + x_i} \\ - \sum_{i=1}^n \frac{1}{\beta + x_i} & \frac{-n \cdot \left[\left[\Gamma(\rho) \cdot \frac{d^2}{d\rho^2} \Gamma(\rho) - \left(\frac{d}{d\rho} \Gamma(\rho) \right)^2 \right] \right]}{\Gamma(\rho)^2} \end{bmatrix}$$

$$I = \begin{bmatrix} -0.856 & -2.242 \\ -2.242 & -7.459 \end{bmatrix}$$

To construct the Wald Statistic we need the variance of ρ , obtained from the information matrix as

$$\text{Varp} := (-I^{-1})_{2,2}$$

$$\text{Varp} = 0.631$$

$$W := (\rho - 1) \cdot (\text{Varp})^{-1} \cdot (\rho - 1)$$

$$W = 7.336$$

This is a large χ^2 , so reject the null.

9.c. The Lagrange Multiplier Test

The restricted estimates of the unknown parameters are

$$\beta := 15.601$$

$$\rho := 1$$

The first order conditions for the maximum of the log likelihood function, evaluated at the restricted estimates are

$$-\rho \cdot \sum_{i=1}^n \frac{1}{\beta + x_i} + \left[\sum_{i=1}^n \left[\frac{y_i}{(\beta + x_i)^2} \right] \right] = 3.742 \cdot 10^{-5}$$

$$\left(- \sum_{i=1}^{20} \ln(\beta + x_i) - n \cdot \text{Psi}(\rho) \right) + \sum_{i=1}^{20} \ln(y_i) =$$

The information matrix evaluated at the restricted estimates

$$I := \begin{bmatrix} \rho \cdot \sum_{i=1}^n \frac{1}{(\beta + x_i)^2} - 2 \cdot \sum_{i=1}^n \frac{y_i}{(\beta + x_i)^3} & - \sum_{i=1}^n \frac{1}{\beta + x_i} \\ - \sum_{i=1}^n \frac{1}{\beta + x_i} & -n \cdot \left[\frac{\Gamma(\rho) \cdot \frac{d^2}{d\rho^2} \Gamma(\rho) - \left(\frac{d}{d\rho} \Gamma(\rho) \right)^2}{\Gamma(\rho)^2} \right] \end{bmatrix}$$

$$I = \begin{bmatrix} -0.022 & -0.669 \\ -0.669 & -32.899 \end{bmatrix}$$

$$LM := \begin{bmatrix} 3.742 \cdot 10^{-5} & 7.916 \end{bmatrix} \cdot (-I)^{-1} \cdot \begin{bmatrix} 3.742 \cdot 10^{-5} \\ 7.916 \end{bmatrix}$$

$$LM = 5.116$$

Although the smallest of the three realized test statistics, this is still above even χ^2 at the 2.5% level of test.